1 General linear hypothesis

- the linear multiple linear regression model is $y = X\beta + \epsilon$.
- the general linear hypothesis is $H_0: T\beta = c$, vs $H_A: T\beta \neq c$.
- where T is an m by p matrix having r independent rows, and c is a $m \times 1$ vector of constants.
- The test statistic is given by

$$F = \frac{(T\hat{\beta} - c)' \left[T(X'X)^{-1}T'\right]^{-1} (T\hat{\beta} - c)/r}{SSE_{full}/(n-p)}$$

• with p-value $P(F_{r,n-p} > F_{obs})$.

2 a general example

Suppose $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \epsilon$.

- to test $\beta_2 = \beta_3$, use T = [0, 0, 1, -1] with c = 0.
- to test $\beta_2 = \beta_3$ and $\beta_1 = 2$, use a matrix T with two rows and four columns. First row of T is [0, 0, 1, -1] and second row of T is [0, 1, 0, 0]. c is [0, 2]'.

-T has two linearly independent rows.

– the F statistic will have 2 numerator, and n-4 denominator degrees of freedom.

3 Application to oneway analysis of variance

Here is a slightly more complicated example. In the one way analysis of variance example with four diets A, B, C, and D, consider the model

$$Y = \beta_1 I_A + \beta_2 I_B + \beta_3 I_C + \beta_4 I_D + \epsilon$$

Here we have dropped the intercept and used all 4 indicator variables.

Note that the mean of Y is β_1 for diet A, β_2 for diet B, β_3 for diet C, and β_4 for diet D. In this respect, this form of the model is really the most natural way to write the oneway ANOVA model using multiple regression.

Consider the hypothesis H_0 : $\beta_1 = \beta_2, \beta_2 = \beta_3, \beta_3 = \beta_4$. This is clearly equivalent to the usual ANOVA hypothesis H_0 : $\mu_A = \mu_B = \mu_C = \mu_D$.

where T is a 3X4 matrix with first row (1,-1,0,0), 2nd row (0,1,-1,0), 3rd row (0,0,1,-1), the hypothesis is equivalent to $H_0: T\beta = 0$, which is in the form of the general linear hypothesis.

Clearly there are many matrices T which give an equivalent to the usual ANOVA hypothesis. The following evaluates the general linear hypothesis test statistic F.

```
> weight=c(3.42 ,3.17 ,3.34 ,3.65 ,
+ 3.96 ,3.63 ,3.72 ,3.93 ,
+ 3.87 ,3.38 ,3.81 ,3.77 ,
+ 4.19 ,3.47 ,3.66 ,4.18 ,
+ 3.58 ,3.39 ,3.55 ,4.21 ,
+ 3.76 ,3.41 ,3.51 ,3.88 ,
+ 3.84 ,3.55
                   ,3.96,
                   ,3.91)
+
       3.44
  diet=as.factor(c(rep(c("A", "B", "C", "D"), 6), "A", "B", "D", "B", "D"))
>
  IA=ifelse(diet=="A",1,0)
>
>
  IB=ifelse(diet=="B",1,0)
 IC=ifelse(diet=="C",1,0)
>
> ID=ifelse(diet=="D",1,0)
>
 X=cbind(IA,IB,IC,ID)
> bhat=solve(t(X)%*%X)%*%t(X)%*%weight
  fits=X%*%bhat
>
> resids=weight-fits
 SSE=sum(resids<sup>2</sup>)
>
> dfe=length(weight)-dim(X)[2]
> MSE=SSE/dfe
  sqrt(MSE) # RESIDUAL STANDARD ERROR
>
[1] 0.1892526
> #fit a model include all 4 indicators, but no intercept
> lm.out= lm(weight~IA+IB+IC+ID-1)
> summary(lm.out) #note the point estimates
```

lm(formula = weight ~ IA + IB + IC + ID - 1)

Call:

```
Residuals:
    Min
               1Q
                   Median
                                ЗQ
                                        Max
-0.38286 -0.05625 -0.00625 0.12000 0.38714
Coefficients:
   Estimate Std. Error t value Pr(>|t|)
IA 3.80286
              0.07153 53.16 <2e-16 ***
IB 3.43000
              0.06691
                        51.26 <2e-16 ***
IC 3.59833
              0.07726 46.57 <2e-16 ***
ID 3.93625
                       58.83 <2e-16 ***
              0.06691
___
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.1893 on 25 degrees of freedom
Multiple R-squared: 0.9977,
                                  Adjusted R-squared:
                                                       0.9974
F-statistic: 2771 on 4 and 25 DF, p-value: < 2.2e-16
>
                  #and that the usual F test is testing the wrong thing
>
> #now carry out F TEST FOR GENERAL LINEAR HYPOTHESIS
> T=matrix(c(1,-1,0,0,
+
             0, 1, -1, 0,
+
             0,0,1,-1),byrow=T,ncol=4); T
     [,1] [,2] [,3] [,4]
[1,]
        1
           -1
                 0
                      0
[2,]
            1
                -1
                       0
        0
[3,]
            0
               1
                     -1
        0
> b=c(0,0,0)
> numSS=t(T%*%bhat-b)%*%solve(T%*%solve(t(X)%*%X)%*%t(T))%*%(T%*%bhat-b)
> numeratorMS=numSS/dim(T)[1]
> Fobs=numeratorMS/MSE
> c(numSS,numeratorMS,Fobs)
[1] 1.1649036 0.3883012 10.8413906
> #CHECK
> summary(lm(weight~IA+IC+ID))
Call:
lm(formula = weight ~ IA + IC + ID)
Residuals:
               1Q
                   Median
                                ЗQ
                                        Max
     Min
-0.38286 -0.05625 -0.00625 0.12000 0.38714
```

Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 3.43000 0.06691 51.262 < 2e-16 *** IA 0.37286 0.09795 3.807 0.000813 *** IC 0.16833 0.10221 1.647 0.112079 ID 0.50625 0.09463 5.350 1.51e-05 *** ----Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 0.1893 on 25 degrees of freedom Multiple R-squared: 0.5654, Adjusted R-squared: 0.5132 F-statistic: 10.84 on 3 and 25 DF, p-value: 9.502e-05